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The new Stata command *spxtsfa* can be downloaded from Github
(<https://github.com/kerrydu/spxtsfa>).

Fitting spatial stochastic frontier models in Stata

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Abstract. In this article, we introduce a new command `spxtsfa` for fitting spatial stochastic frontier models in Stata. Over the last decades, an important theoretical progress of stochastic frontier models is the incorporation of various types of spatial components. Models with the ability to account for spatial dependence and spillovers have been developed for efficiency and productivity analysis, drawing extensive attention from industry and academia. Due to the unavailability of the statistical packages, the empirical applications of the new stochastic frontier models appear to be lagging. The `spxtsfa` command provides a routine for estimating the spatial stochastic frontier models in the style of [Orea and Álvarez \(2019\)](#) and [Galli \(2022\)](#), enabling users to handle different sources of spatial dependence. In the presented article, we introduce the spatial stochastic frontier models, describe the syntax and options of the new command, and provide several examples to illustrate its usage.

Keywords: stochastic frontier models, SFA, spatial dependence, technical efficiency, spillovers

1 Introduction

Producers might fail in optimizing their production activities, causing deviation from the maximum output or the minimum cost. Economic researchers proposed the concept of technical efficiency, which measures how well a producer is utilizing its resources to produce goods or services. A technically efficient organization makes the maximum outputs given the amount of inputs or uses the minimum amount of inputs to produce a given level of output. On the contrary, technically inefficient organization produce fewer outputs given the same inputs or uses more inputs than necessary to produce the same output. Technical efficiency is important because it allows organizations or economies to achieve their goals with the least amount of resources possible, which can

lead to cost savings and increased profitability.

Aigner et al. (1977) and Meeusen and van Den Broeck (1977) introduced stochastic frontier models for evaluating technical efficiency. The essential concept behind these models is to divide the observed output of a production process into two components, namely the "frontier" output, signifying the maximum feasible output, given the inputs utilized in the production process, and the "residual" output, denoting the production process's inefficiency. Following these initial works, stochastic frontier models gained extensive use as a tool for scrutinizing productivity and efficiency.

Methodologically, econometricians have expanded the horizons of stochastic frontier models in various directions. To name a few, Battese and Coelli (1995) incorporated the determinants of inefficiency. Wang (2003) developed the stochastic frontier model with scaling properties to capture the shape of the distribution of inefficiency. Greene (2005) extended the stochastic models with the random effects and the "true" fixed effects. Belotti and Ilardi (2018), Chen et al. (2014), and Wang and Ho (2010) circumvented the "incidental parameter problem" in the fixed effects stochastic frontier model through model transformation. Karakaplan and Kutlu (2017) developed an endogenous stochastic frontier model to control for the endogeneity in the frontier or inefficiency.

In recent years, stochastic frontier models have undergone further extension to account for spatial dependence and spatial spillover effects. Glass et al. (2016) constructed a spatial Durbin stochastic model considering both global and local spatial dependence. Kutlu et al. (2020) proposed a spatial stochastic frontier model with endogenous frontier and environmental variables. Glass et al. (2016) and Kutlu et al. (2020) combine the concepts of spatial econometrics and stochastic frontier analysis by including the spatial lag of the dependent variable. On the other hand, Orea and Álvarez (2019) developed a new stochastic frontier model with spatial correlation in both noise and inefficiency terms. Galli (2022) integrated the two different modeling ideas to specify four different sources of spatial dependence fully.

With the increasing demand in the last decades to analyze technical efficiency, Stata provides official commands `frontier` and `xtfrontier` for cross-sectional and panel stochastic model estimation, respectively. Belotti et al. (2013) developed `sfcross` and `sfpnl` commands accommodating more different distribution assumptions and allowing fixed-effect and random-effect models with the consideration of heteroscedasticity. Karakaplan (2017) introduced the `sfkk` command for estimating endogenous stochastic frontier models. Karakaplan (2018) supplemented the `xtsfkk` command for fitting the endogenous panel stochastic frontier model. Kumbhakar et al. (2015) provides a practitioner's guide to stochastic frontier analysis with a suite of Stata commands (including `sfmodel`, `sfpnl`, `sf_fixeff`, and `sfpnl`).

In this article, we introduce `spxtsfa`, a new command for fitting spatial stochastic frontier models in the style of Orea and Álvarez (2019) and Galli (2022). The proposed `spxtsfa` command not only allows getting more accurate inefficiency scores (see e.g. Orea et al. 2018) but also examining relevant economic issues that a non-spatial stochastic frontier model tends to overlook.

For instance, in microdata applications, the new command can be used to test whether the production/cost function can be viewed as a purely deterministic (engineering) process where the firm controls all the inputs (see e.g. [Druska and Horrace 2004](#)). A distinctive feature of the `spxtsfa` command is that it allows estimating a stochastic frontier model with cross-sectional correlation in the inefficiency term, a specification that is useful in applications where some firms benefit from best practices implemented in adjacent firms due to, for instance, agglomeration economies, knowledge spillovers, technology diffusion or R&D spillovers. This could especially be the case if (local) firms belong to communitarian networks (e.g. cooperatives) or common technicians (consultants) are advising all local firms. In practice, the proposed `spxtsfa` command can be useful to capture a kind of behavioral correlation, for instance when firms tend to “keep an eye” on the decisions of other peer firms trying to overcome the limitations caused by the lack of information or they simply emulate each other. It is finally germane to mention that the `spxtsfa` command also allows capturing cross-sectional effects that might be caused by non-spatial factors (e.g., the regulation environment) if we define appropriately the so-called weight (W) matrix. A proper definition of the W matrix might, for instance, allow us to examine the existence of knowledge spillovers from supplier and user firms.

As [Orea and Álvarez \(2019\)](#) point out, the proposed `spxtsfa` command can be implemented using macro-level data (e.g. data of countries, regions or industries) due to the abundant evidence of important feedback processes between neighboring or non-distant regions justify the use of SAR and Durbin frontier functions in macrodata applications. The spatial weight matrix specification commonly adopted in regional economics is based on geographical distance. However, as aforementioned, the weight matrix can be defined using a non-spatial criterion. In this sense, [Liu and Sickles \(2023\)](#) state that the mode of production in the world economy is characterized by the division of global value chains (GVCs) and, hence, the spatial weight matrix should be constructed using the economic distance between industries within/across national economies. In this case, the proposed `spxtsfa` command can be used to estimate spatial SAR and Durbin frontier functions in order to examine the diffusion of knowledge and technology among the participants in the international production network. It is also makes sense to estimate a stochastic frontier model with cross-sectional correlation in the inefficiency term using macrodata if we change the interpretation of the estimated correlation. In these applications, the spatial correlation in the inefficiency term likely captures barriers and distortions to the efficient allocation of resources across firms that are common to several regions, such as regulation, labor market trends or common institutions (see e.g. [Orea et al. 2023](#)).

The remainder of this article unfolds as follows: Section 2 provides a brief description of the models in [Orea and Álvarez \(2019\)](#) and [Galli \(2022\)](#); Section 3 explain the syntax and options of `spxtsfa`; Section 4 and 5 present simulated data examples to illustrate the usage of the command; and section 6 concludes the article.

2 The model

In this section, we briefly describe the spatial stochastic frontier models developed by [Orea and Álvarez \(2019\)](#) and [Galli \(2022\)](#). The exposition here is only introductory. Please refer to the cited papers for more technical details.

Based on the transposed version of [Wang and Ho \(2010\)](#) model, [Orea and Álvarez \(2019\)](#) proposed a spatial stochastic frontier model which accommodates spatially-correlated inefficiency and noise terms. The model is formulated as in Eqs.(1)-(3), for $i = 1, \dots, N$ and $t = 1, \dots, T$:

$$Y_{it} = X_{it}'\beta + \tilde{v}_{it} - s\tilde{u}_{it} \quad (1)$$

$$\tilde{v}_{it} = v_{it} + \gamma W_i^{vt} \tilde{v}_{.t} \quad (2)$$

$$\tilde{u}_{it} = u_{it} + \tau W_i^{ut} \tilde{u}_{.t} \quad (3)$$

Eq.(1) describes the stochastic frontier function where Y_{it} is the dependent variable and X_{it} is a $k \times 1$ vector of variables shaping the frontier; $s = 1$ for the production function and $s = -1$ for the cost function; \tilde{v}_{it} and \tilde{u}_{it} represent idiosyncratic noise and inefficiency, respectively. In Eqs.(2) and (3), $W_i^{vt} = (W_{i1}^{vt}, \dots, W_{iN}^{vt})$ and $W_i^{vt} = (W_{i1}^{vt}, \dots, W_{iN}^{vt})$ are two known $1 \times N$ cross-sectional weight vectors depicting the structure of the cross-sectional relationship for idiosyncratic noise and inefficiency terms, respectively; $\tilde{v}_{.t} = (\tilde{v}_{1t}, \dots, \tilde{v}_{Nt})'$ and $\tilde{u}_{.t} = (\tilde{u}_{1t}, \dots, \tilde{u}_{Nt})'$; v_{it} is a random variable following the distribution $N(0, \sigma_v^2)$ and $u_{it} = h(Z_{it}'\delta)u_t^*$. $h(Z_{it}'\delta)$ is the scaling function where Z_{it} is a $l \times 1$ vector of variables affecting individuals' inefficiency and u_t^* is a non-negative random variable following the distribution $N^+(0, \sigma_u^2)$. Using matrix notation, we can rewrite Eqs.(2) and (3) as

$$\tilde{v}_{.t} = (I_N - \gamma W^{vt})^{-1} v_{.t} \quad (4)$$

$$\tilde{u}_{.t} = (I_N - \tau W^{ut})^{-1} h(Z_{.t}\delta)u_t^* = \tilde{h}_{.t}u_t^* \quad (5)$$

where $Z_{.t} = (Z_{1t}, \dots, Z_{Nt})'$; $\tilde{h}_{.t} = (I_N - \tau W^{ut})^{-1} h(Z_{.t}\delta)$.

The above model captures the spatial correlation of the random error and inefficiency terms with the spatial autoregressive (SAR) process ¹. Referring to [Wang and Ho \(2010\)](#), we can obtain the following log-likelihood function for each period t :

$$\begin{aligned} \ln L_t = & -\frac{N}{2} \ln(2\pi) - \frac{1}{2} \ln |\Pi| - \frac{1}{2} \tilde{\varepsilon}_{.t}' \Pi^{-1} \tilde{\varepsilon}_{.t} \\ & + \frac{1}{2} \left(\frac{\mu_*^2}{\sigma_*^2} \right) + \ln \left[\sigma_* \Phi \left(\frac{\mu_*}{\sigma_*} \right) \right] - \ln \left(\frac{1}{2} \sigma_u \right) \end{aligned} \quad (6)$$

¹[Orea and Álvarez \(2019\)](#) also considered a specification of the spatial moving average process.

where $\Pi = \sigma_v^2(I_N - \rho W^{yt})^{-1}[(I_N - \rho W^{yt})^{-1}]'$; $\tilde{\varepsilon}_{.t} = (\tilde{\varepsilon}_{1t}, \dots, \tilde{\varepsilon}_{Nt})'$, $\tilde{\varepsilon}_{it} = s(Y_{it} - X'_{it}\beta)$, and

$$\mu_* = \frac{-\tilde{\varepsilon}'_{.t}\Pi^{-1}\tilde{h}_{.t}}{\tilde{h}'_{.t}\Pi^{-1}\tilde{h}_{.t} + 1/\sigma_u^2} \quad (7)$$

$$\sigma_*^2 = \frac{1}{\tilde{h}'_{.t}\Pi^{-1}\tilde{h}_{.t} + 1/\sigma_u^2} \quad (8)$$

Galli (2022) further incorporated the spatial lags of the dependent variable and the input variables into Orea and Álvarez (2019) model, which additionally measures global and local spatial spillovers affecting the frontier function. The model is expressed as

$$Y_{it} = \rho W_i^{yt} Y_{.t} + X'_{it}\beta + W_i^{xt} X_{.t}\theta + \tilde{v}_{it} + s\tilde{u}_{it} \quad (9)$$

where $W_i^{yt} = (W_{i1}^{yt}, \dots, W_{iN}^{yt})$ and $W_i^{xt} = (W_{i1}^{xt}, \dots, W_{iN}^{xt})$ are two known $1 \times N$ cross-sectional weight vectors²; $Y_{.t} = (Y_{1t}, \dots, Y_{Nt})'$; $X_{.t} = (X_{1t}, \dots, X_{Nt})'$. This model gives rise to the following log-likelihood function for each period t :

$$\begin{aligned} \ln L_t = & \ln |I_N - \rho W^{yt}| - \frac{N}{2} \ln(2\pi) - \frac{1}{2} \ln |\Pi| - \frac{1}{2} \tilde{\varepsilon}'_{.t} \Pi^{-1} \tilde{\varepsilon}_{.t} \\ & + \frac{1}{2} \left(\frac{\mu_*^2}{\sigma_*^2} \right) + \ln \left[\sigma_* \Phi \left(\frac{\mu_*}{\sigma_*} \right) \right] - \ln \left(\frac{1}{2} \sigma_u \right) \end{aligned} \quad (10)$$

where $\tilde{\varepsilon}_{.t} = (\tilde{\varepsilon}_{1t}, \dots, \tilde{\varepsilon}_{Nt})'$, $\tilde{\varepsilon}_{it} = s(Y_{it} - X'_{it}\beta - \rho W_i^{yt} Y_{.t} - W_i^{xt} X_{.t}\theta)$.

Summing the time-specific log-likelihood functions over all periods yields the overall likelihood function for the whole sample, i.e., $\ln L = \sum_{t=1}^T \ln L_t$. Then, numerically maximize the overall log-likelihood function to obtain consistent estimates of the parameters in the above models. Specifically, we use Stata `ml` `model` routine with the `method-d0` evaluator to program the `spxtsf` command. Following Gude et al. (2018), we parameterize ρ , γ , and τ as Eq.(11) to ensure the standard regularity condition for the spatial autoregressive models.

$$\begin{aligned} \eta &= \left(\frac{1}{r_{\min}} \right) (1 - p) + \left(\frac{1}{r_{\max}} \right) p \\ 0 \leq p &= \frac{\exp(\delta_0)}{1 + \exp(\delta_0)} \leq 1 \end{aligned} \quad (11)$$

where η stands for one of ρ , γ , and τ ; r_{\min} and r_{\max} are respectively the minimum and maximum eigenvalues of the corresponding spatial weight matrix.

In summary, Galli (2022) provided a fully comprehensive specification of four different types of spatial dependence: global spillovers of dependent variable Y_{it} , local spillovers of input variables X_{it} , cross-sectional correlation of idiosyncratic noise v_{it} and inefficiency u_{it} . We term this full model "yxwv-SAR". Some restrictions can be imposed on the specific parameters to generate

²We index W_i^{yt} , W_i^{xt} , W_i^{ut} , and W_i^{vt} with superscript yt , xt , ut , and vt , respectively. This indicates the spatial weight matrix can be time-varying and different across various spatial components

the following models (summarized in Table 1), which can be estimated by the `spxtsfa` command.

Table 1: Specific models with restricted parameters

	yw	xw	yv	yu	y	xw	xv	xu	uv	u	v
ρ		0				0	0	0	0	0	0
θ	0		0	0	0				0	0	0
γ				0	0			0		0	
τ			0		0		0				0

3 The `spxtsfa` command

`spxtsfa` estimates spatial stochastic frontier models in the style of [Orea and Álvarez \(2019\)](#) and [Galli \(2022\)](#).

3.1 Syntax

Estimation syntax

```
spxtsfa depvar [indepvars] uhet(varlist) [ noconstant cost wy(wyspec)
  wx(wxspec) wu(wuspec) wv(wvspec) normalize(norm_method)
  wxvars(varlist) initial(matname) mlmodel(model_options)
  mlsearch(search_options) mlplot mlmax(maximize_options) nolog
  mldisplay(display_options) level(#) lndetmc(numlist)
  te(newvarname) genwxvars delmissing constraints(constraints) ]
```

Version syntax

```
spxtsfa , version
```

Replay syntax

```
spxtsfa [ , level(#) ]
```

3.2 Options

`uhet`(*varlist*) specifies explanatory variables for technical inefficiency variance function depending on a linear combination of *varlist*. It is required.

`noconstant` suppresses constant term.

`cost` specifies the frontier as a cost function. By default, the production function is assumed.

`wy`(*wyspec*) specifies the spatial weight matrix for lagged dependent variable.

The expression is `wy(W1 [W2...WT] [mata array])`. By default, the weight matrices are `Sp` objects. `mata` declares weight matrices are `mata` matrices.

If one weight matrix is specified, it assumes a time-constant weight matrix.

For time-varying cases, T weight matrices should be specified in time order. Alternatively, using array to declare weight matrices are stored in an array. If only one matrix is stored in the specified array, the time-constant weight matrix is assumed. Otherwise, the keys of the array specify time information, and the values store time-specific weight matrices.

`wx(wspec)` specifies the spatial weight matrix for lagged independent variable. The expression is the same as `wy(wspec)`.

`wu(wuspec)` specifies the spatial weight matrix for lagged independent variable. The expression is the same as `wy(wspec)`.

`wv(wvspec)` specifies the spatial weight matrix for lagged independent variable. The expression is the same as `wy(wspec)`.

`normalize(norm_method)` specifies one of the four available normalization techniques: row, col, minmax, and spectral.

`wxvars(varlist)` specifies spatially lagged independent variables.

`initial(matname)` specifies the initial values of the estimated parameters with matrix *matname*.

`mlmodel(model_options)` specifies the `ml model` options.

`mlsearch(search_options)` specifies the `ml search` options.

`mlplot` specifies using `ml plot` to search better initial values of spatial dependence parameters.

`mlmax(maximize_options)` specifies the `ml maximize` options.

`nolog` suppresses the display of the criterion function iteration log.

`mldisplay(display_options)` specifies the `ml display` options.

`level(#)` sets confidence level; default is level(95).

`lndetmc(numlist)` uses the trick of [Barry and Kelley Pace \(1999\)](#) to solve the inverse of $(I_N - \rho W)$. The order of *numlist* is iterations, maxorder. `lndetmc(50 100)` specifies that the number of iterations is 50 and the maximum order of moments is 100.

`te(newvarname)` specifies a new variable name to store the estimates of technical efficiency.

`genwxvars` generates the spatial Durbin terms. It is activated only when `wxvars(varlist)` is specified.

`delmissing` allows estimation when missing values are present by removing the corresponding units from spatial matrix.

`constraints(constraints)` specifies linear constraints for the estimated model.

3.3 Dependency of spxtsfa

`spxtsfa` depends on the *moremata* package. If not already installed, you can install it by typing `ssc install moremata`.

4 Examples

In this section, we use simulated data to exemplify the use of the *spxtsfa* command. Referring to , we first consider the *yxuv*-SAR model specified by the following data-generating process (DGP 1) with $i = 1, \dots, 300$ and $t = 1, \dots, 20$,

$$Y_{it} = 0.3W_iY_{i,t} + 2X_{it} + 0.3W_iX_{i,t} + \tilde{v}_{it} - \tilde{u}_{it} \quad (12)$$

where \tilde{v}_{it} and \tilde{u}_{it} are defined as in Eqs.(2) and (3) with $\gamma = 0.3$, $\tau = 0.3$, $\delta = 2$, $\sigma_u^2 = 0.2$ and $\sigma_v^2 = 0.2$. All the spatial matrices for the four spatial components are the same and time-invariant, created from a binary contiguity spatial weight matrix. We generate the exogenous variables X_{it} and Z_{it} from the standard normal distribution, respectively. With the sample generated by DGP 1, we can fit the model in the following syntax.

```
. use spxsfafa_DGP1.dta
. xtset id t
Panel variable: id (strongly balanced)
Time variable: t, 1 to 20
Delta: 1 unit

. * importing spatial weight matrix from spxsfafa_wmat1.mmat
. mata mata matuse spxsfafa_wmat1.mmat,replace
(loading w1[300,300])
. * fitting the model
. spxsfafa y x, uhets(z) noconstant wy(w1,mata) wx(w1,mata) wu(w1,mata) wv(w1,mata) wxvars(x) nolog
Spatial frontier model(yxuv-SAR)
Log likelihood = -1727.016
Number of obs = 6,000
Wald chi2(2) = 118937.24
Prob > chi2 = 0.0000
```

	y	Coefficient	Std. err.	z	P> z	[95% conf. interval]	
frontier							
	x	1.993915	.0065251	305.58	0.000	1.981126	2.006704
	W_x	.4435823	.0373189	11.89	0.000	.3704386	.516726
uhets							
	z	2.000371	.0013412	1491.49	0.000	1.997742	2.002999
/lnsigma2_u							
	/lnsigma2_v	-2.098104	.3163094	-6.63	0.000	-2.718059	-1.478149
Wu							
	_cons	.6605993	.0317043	20.84	0.000	.5984599	.7227386
Wv							
	_cons	.5806681	.0318346	18.24	0.000	.5182735	.6430627
sigma2_u							
	sigma2_v	.1226888	.0388076	3.16	0.002	.0660027	.2280593
	rho	.1944444	.003578	54.34	0.000	.1875567	.2015851
	tau	.3187581	.0142397	22.39	0.000	.2905787	.3463849
	gamma	.282414	.014646	19.28	0.000	.2534626	.3108598
		.2795936	.02392	11.69	0.000	.2320763	.3257792

The output shows that the command fits seven equations with `ml model1`. The frontier equation has two explanatory variables X_{it} and $W_iX_{i,t}$. The scaling function `uhets()` has one explanatory variable Z_{it} . Two equations (`/lnsigma2_u`

and $/\ln\sigma_{2_v}$) are constructed for the variance parameters σ_u^2 and σ_v^2 which are transformed by the function $\exp(\cdot)$. Three Equations (Wy, Wu, and Wv) handle the spatial dependence parameters ρ , τ , and γ , which are parameterized as Eq.(11). We directly include the spatial Durbin term $W_i X_{i,t}$ in the frontier equation (represented as W_x) such that we do not need to fit a separate equation. The bottom of the table reports the transformed parameters in the original metric.

We consider the restricted model uv -SAR with time-varying spatial weight matrices as the second example. The DGP 2 is described as

$$Y_{it} = 1 + 2X_{it} + \tilde{v}_{it} - \tilde{u}_{it}, i = 1, \dots, 300; t = 1, \dots, 10 \quad (13)$$

where the other parameters are set the same as the DGP 1 except for $W_i^{ut} = W_i^{vt} = W_i^t$. The following syntax estimates the model alongside the results.

```
. use spxsfadgp2.dta
. xtset id t
Panel variable: id (strongly balanced)
Time variable: t, 1 to 10
Delta: 1 unit

. * importing spatial weight matrices from spxsfawmat2.mmat
. mata mata matus spxsfawmat2.mmat,replace
(loading w1[300,300], w10[300,300], w2[300,300], w3[300,300], w4[300,300],
w5[300,300], w6[300,300], w7[300,300], w8[300,300], w9[300,300])

. * fitting the model
. local w w1 w2 w3 w4 w5 w6 w7 w8 w9 w10
. spxsfadgp2 y x, uhet(z) wu(`w',mata) wv(`w',mata) te(eficiency) nolog
Spatial frontier model(uv-SAR)                               Number of obs =    3,000
                                                              Wald chi2(1)    = 43686.91
Log likelihood = -1336.482                                   Prob > chi2     =  0.0000
```

	y	Coefficient	Std. err.	z	P> z	[95% conf. interval]	
frontier							
	x	2.015288	.0096419	209.01	0.000	1.99639	2.034186
	_cons	.9415143	.0160786	58.56	0.000	.9100008	.9730278
uhet							
	z	2.000242	.0020671	967.66	0.000	1.99619	2.004293
/lnsigma2_u							
/lnsigma2_v							
		-2.006684	.4473506	-4.49	0.000	-2.883475	-1.129893
		-1.300024	.0260099	-49.98	0.000	-1.351002	-1.249045
Wu							
	_cons	.582383	.0031549	184.59	0.000	.5761995	.5885666
Wv							
	_cons	.5374655	.0601775	8.93	0.000	.4195198	.6554113
	sigma2_u	.1344337	.060139	2.24	0.025	.05594	.3230678
	sigma2_v	.2725253	.0070883	38.45	0.000	.2589806	.2867784
	tau	.2832028	.0014508	195.21	0.000	.2803569	.2860438
	gamma	.262419	.0280135	9.37	0.000	.206716	.3164261

In the second example, we use option `te(efficiency)` to store the estimated efficiency score in a new variable `efficiency`. To show the usage of the `delmissing` option, we replace the first observation of Y_{it} with missing value and re-run the above codes which gives rise to error information *"missing values found. use delmissing to remove the units from the spmatrix"*. The inclusion of the `delmissing` option addresses this issue and the generated variable `__e_sample__` records the regression sample.

```
. * replace the first observation of y with missing value
. replace y=. in 1
(1 real changes made)
. local w w1 w2 w3 w4 w5 w6 w7 w8 w9 w10
. * estimation is aborted
. cap noi spxstfa y x, uhets(z) wu(`w`,mata) wv(`w`,mata) nolog
missing values found. use delmissing to remove the units from the spmatrix
invalid syntax

. * re-estimation with delmissing option
. local w w1 w2 w3 w4 w5 w6 w7 w8 w9 w10
. spxstfa y x, uhets(z) wu(`w`,mata) wv(`w`,mata) delmissing nolog
missing values found. The corresponding units are deleted from the spmatrix

Spatial frontier model(uv-SAR)                                Number of obs =    2,999
Wald chi2(1) = 43688.15
Log likelihood = -1336.2158                                Prob > chi2    =    0.0000
```

	y	Coefficient	Std. err.	z	P> z	[95% conf. interval]	
frontier							
	x	2.01521	.0096414	209.02	0.000	1.996313	2.034106
	_cons	.9409358	.0160983	58.45	0.000	.9093837	.972488
uhet							
	z	2.000244	.0020675	967.48	0.000	1.996192	2.004296
/lnsigma2_u		-2.006689	.4473506	-4.49	0.000	-2.88348	-1.129898
/lnsigma2_v		-1.300032	.0260157	-49.97	0.000	-1.351022	-1.249042
Wu							
	_cons	.5823337	.0031574	184.44	0.000	.5761453	.588522
Wv							
	_cons	.5400995	.0602196	8.97	0.000	.4220712	.6581278
sigma2_u		.134433	.0601387	2.24	0.025	.0559397	.3230661
sigma2_v		.2725231	.0070899	38.44	0.000	.2589755	.2867794
tau		.2831801	.0014519	195.04	0.000	.2803319	.2860233
gamma		.2636448	.0280137	9.41	0.000	.2079367	.3176476

```
Missing values found
The regression sample recorded by variable __e_sample__
```

Finally, we consider another restricted model xuv -SAR with different spatial weight matrices, one of which is time-varying, and the others are time-constant.

The model is described as DGP 3:

$$Y_{it} = 1 + 2X_{it} + 0.5W_i^{xt} + \tilde{v}_{it} + \tilde{u}_{it}, i = 1, \dots, 300; t = 1, \dots, 10 \quad (14)$$

where the other parameters are set the same as the DGP 1 except for $W_i^{ut} = W_i^u$ and $W_i^{vt} = W_i^v$. Different from DGP 1 and DGP 2, which set the production function frontier, DGP 3 specifies a cost function. The estimation of the model is shown as follows.

```
. use spxstfa_DGP3.dta
. xtset id t
Panel variable: id (strongly balanced)
Time variable: t, 1 to 10
Delta: 1 unit

. * importing spatial weight matrices from spxstfa_wmat2.mmat
. mata mata matuse spxstfa_wmat2.mmat,replace
(loading w1[300,300], w10[300,300], w2[300,300], w3[300,300], w4[300,300],
w5[300,300], w6[300,300], w7[300,300], w8[300,300], w9[300,300])
. * fitting the model
. local w w1 w2 w3 w4 w5 w6 w7 w8 w9 w10
. mat b = (1,1,1,1,-1,-1,0.5,0.5)
. spxstfa y x, uhets(z) wu(w2,mata) wv(w1,mata) wxvars(x) ///
> wx(`w',mata) cost init(b) genwxvars nolog

Spatial frontier model(xuv-SAR)                                Number of obs =    3,000
                                                                Wald chi2(2)    = 57430.99
Log likelihood = -872.06794                                     Prob > chi2     =  0.0000
```

	y	Coefficient	Std. err.	z	P> z	[95% conf. interval]	
frontier							
	x	1.995734	.0083857	237.99	0.000	1.979298	2.012169
	W_x	.5065867	.0221169	22.90	0.000	.4632384	.549935
	_cons	.9916378	.0126993	78.09	0.000	.9667476	1.016528
uhets							
	z	1.99976	.0010759	1858.74	0.000	1.997652	2.001869
/lnsigma2_u							
	/lnsigma2_u	-1.699017	.4472788	-3.80	0.000	-2.575668	-.822367
/lnsigma2_v							
	/lnsigma2_v	-1.615448	.0260424	-62.03	0.000	-1.66649	-1.564406
Wu							
	_cons	.6214083	.0008566	725.44	0.000	.6197294	.6230872
Wv							
	_cons	.6001869	.0595911	10.07	0.000	.4833905	.7169833
sigma2_u							
	sigma2_u	.1828631	.0817908	2.24	0.025	.076103	.4393904
sigma2_v							
	sigma2_v	.1988017	.0051773	38.40	0.000	.188909	.2092123
tau							
	tau	.3010474	.0003894	773.04	0.000	.300284	.3018105
gamma							
	gamma	.291369	.0272628	10.69	0.000	.2370726	.3438504

In the third example, we use `cost` option to specify the type of frontier. The matrix `b` is used as the initial value for the maximum likelihood estimation. The likelihood function of spatial stochastic frontier models is complicated, and

generally difficult to obtain the optimal global solutions. Thus, good initial values would be helpful for fitting spatial stochastic models. Practitioners might fit the non-spatial stochastic models using `fronteir` and `sfpanel` commands to obtain the initial values of the parameters involved in the frontier and the scaling function and then use the `mlplot` option to search initial values for spatially-correlated parameters.

5 Conclusion

Geospatial units are not isolated or separated but connected. For example, the economic trade, social activities, and cultural exchange between different regions affect each other. Such spatial interdependence challenges the traditional econometric methods, which generally assume cross-sectional independence. Spatial econometrics is developed to handle spatial correlation. Recently, researchers combined stochastic frontier models with spatial econometrics to account for various types of spatial effects in the field of efficiency and productivity analysis (Galli 2022; Orea and Álvarez 2019). This article presented a community-contributed command for fitting spatial stochastic frontier models with different sources of spatial dependence. We hope the developed command can provide some convenience to practitioners and reduce the difficulty of model applications, thereby promoting sound empirical research.

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